

Time Series Toolbox

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Outline: Time Series in Practice

- ▶ When and why do we need time series models?
- ▶ Basic models and definitions: white noise, AR1, MA, random walk, stationarity.
- ▶ 3 approaches to time series modelling: ARIMA, Regression, Structural time series / state-space models

Aim today: understand basic difficulties with time series, construct a few simple but useful models

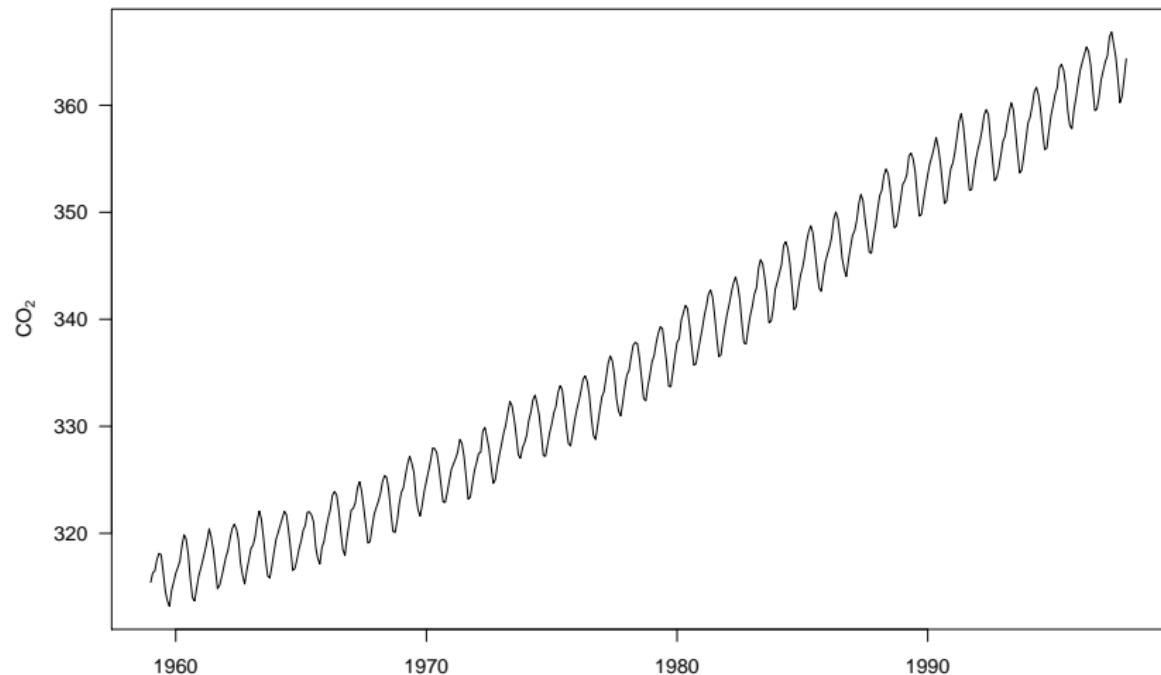
References

- ▶ Hyndman, R.J., & Athanasopoulos, G. (2018) Forecasting: Principles and Practice, 2nd edition, OTexts: Melbourne, Australia. <https://otexts.com/fpp2>
- ▶ Applied Time Series Analysis for Fisheries and Environmental Sciences. E. E. Holmes, M. D. Scheuerell, and E. J. Ward. (2019).
<https://nwfsc-timeseries.github.io/atsa-labs/index.html>

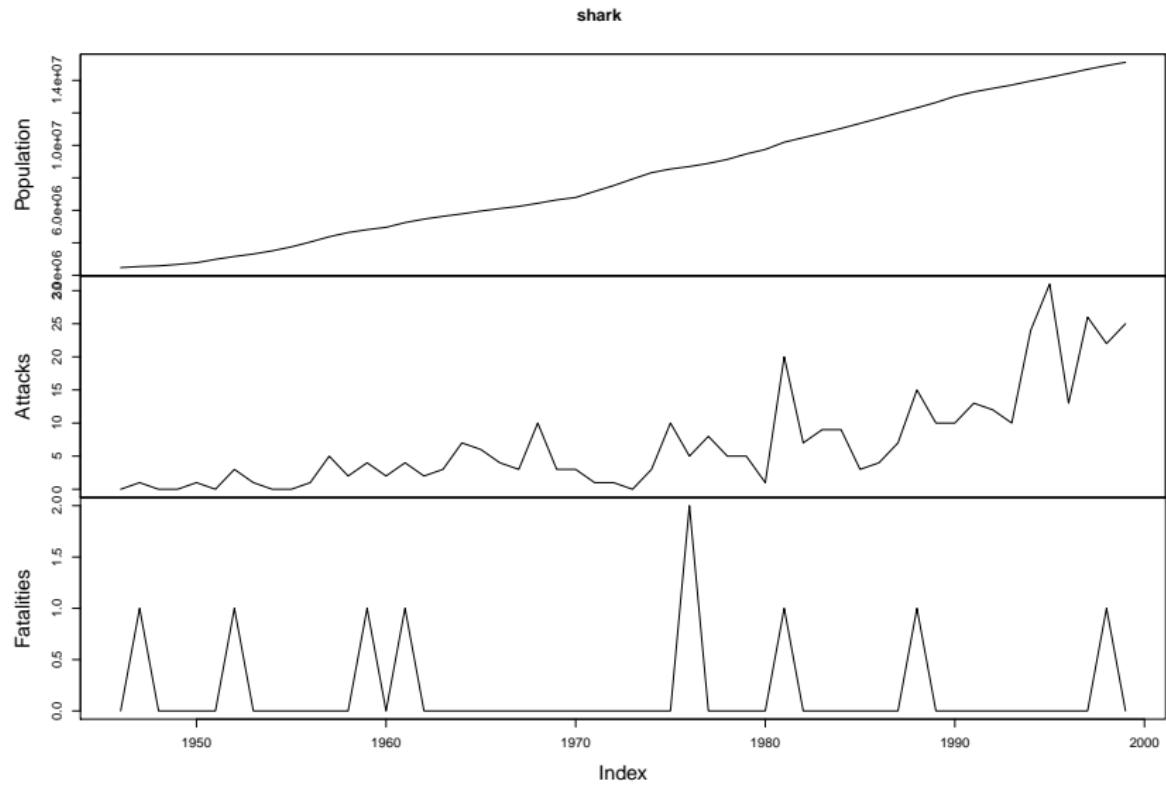
More technical:

- ▶ Shumway, R.H. and Stoffer, D.S., 2017. Time series analysis and its applications: with R examples. Springer.
<https://www.stat.pitt.edu/stoffer/tsa4/>

Motivating Example: Mauna Loa Atmospheric CO₂ Concentration



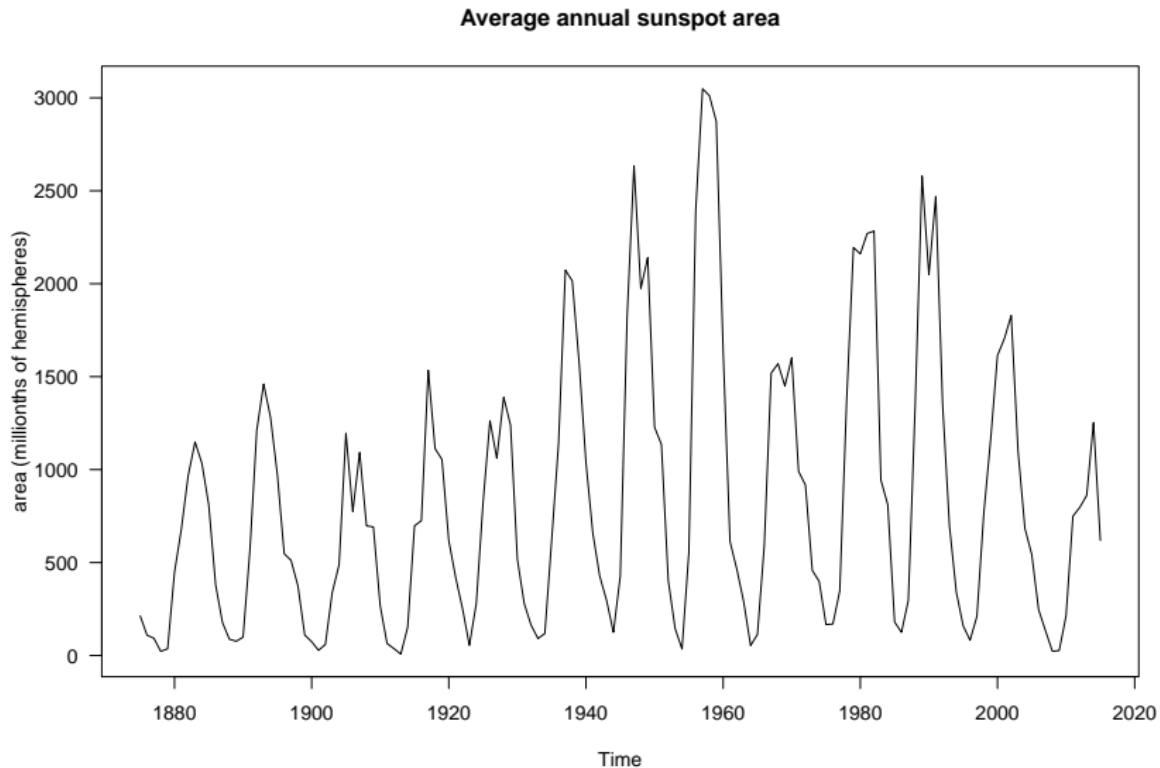
Shark Attacks in Florida



Financial

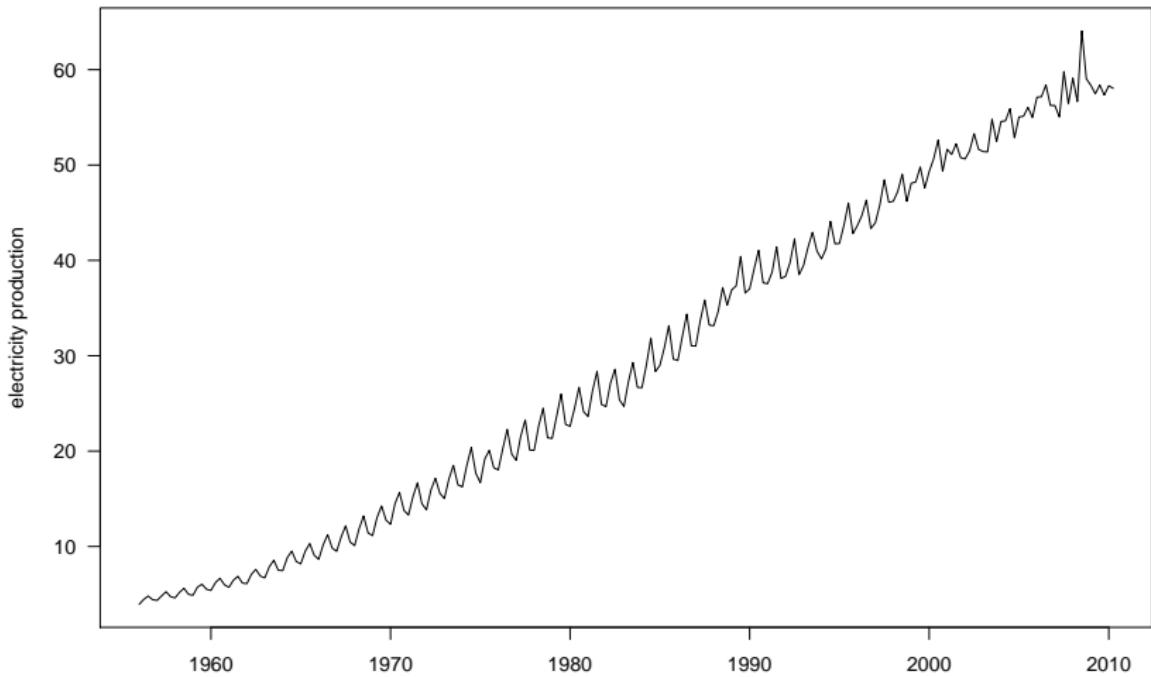


Sunspot Area



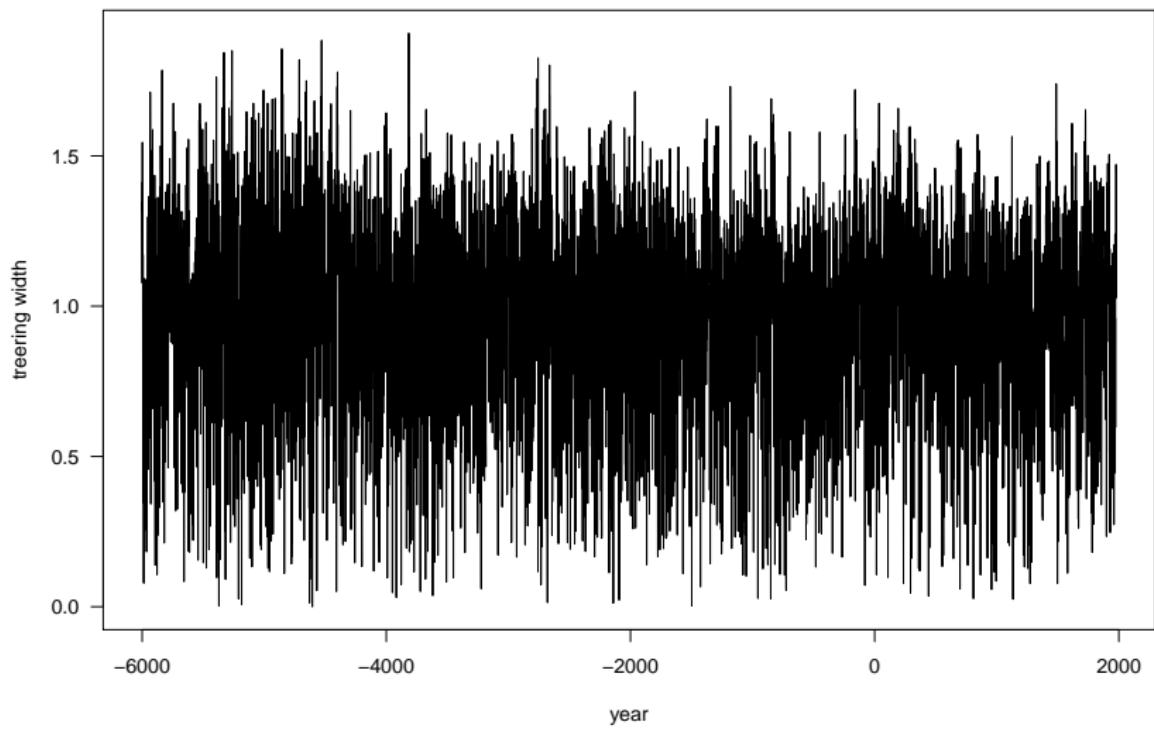
Electricity Production

Quarterly Australian Electricity production



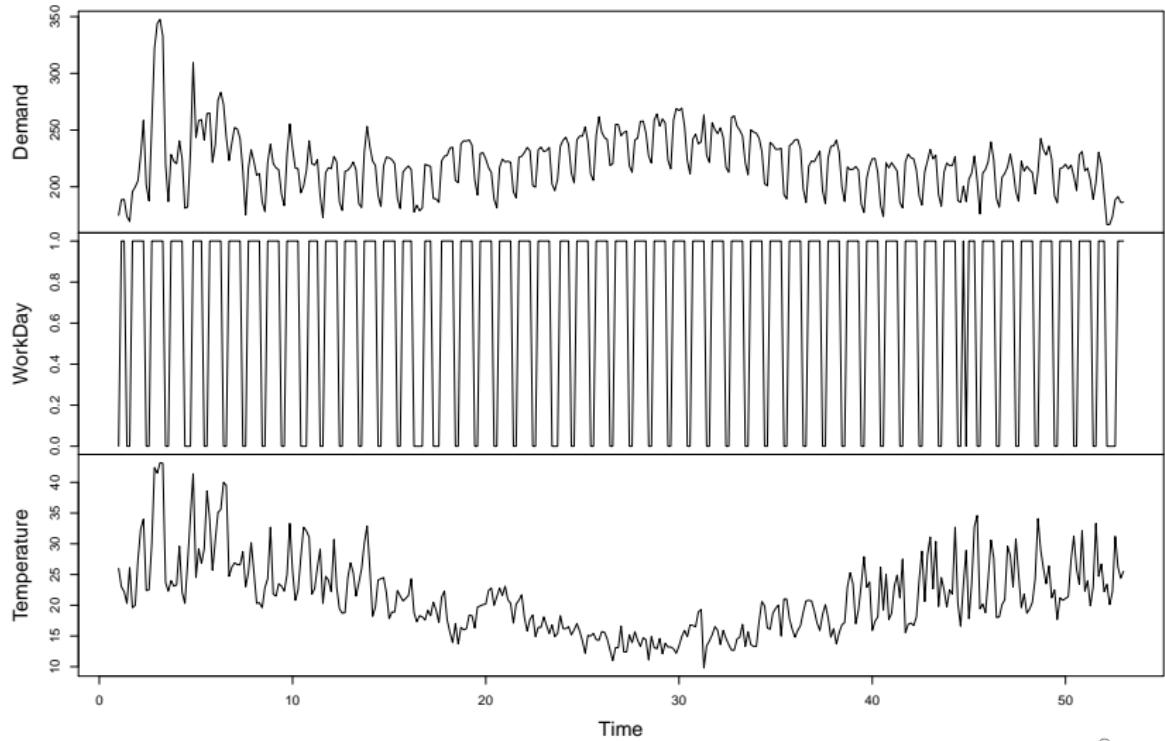
Treerings

Yearly Treering Data, -6000...1979



Electricity Demand

Daily electricity demand, temperature, Victoria, Australia, 2014



Goals of Time Series Analysis

- ▶ prediction / forecast
- ▶ impact of single event
- ▶ study causal patterns

Time Series in R

```
## CO2 data

##          Jan      Feb      Mar      Apr      May      Jun      Jul
## 1959 315.42 316.31 316.50 317.56 318.13 318.00 316.39 31
## 1960 316.27 316.81 317.42 318.87 319.87 319.43 318.01 31
##          Nov      Dec
## 1959 314.66 315.43
## 1960 314.84 316.03

##
##  treeering data

## Time Series:
## Start = -6000
## End = -5981
## Frequency = 1
## [1] 1.345 1.077 1.545 1.319 1.413 1.069 0.489 1.171 0.8
## [12] 0.846 0.837 0.079 0.829 0.919 0.776 0.081 0.876 0.2
```

Time Series in R

Create time series of quarterly data:

```
ts(rnorm(40), frequency = 4, start = c(2019, 2))
```

##	Qtr1	Qtr2	Qtr3	Qtr4
## 2019	0.33524522	-0.71583544	-0.13751188	
## 2020	1.03208627	1.26466849	0.88699606	0.26858419
## 2021	1.10018262	-0.58865896	0.14023571	-0.27432658
## 2022	-0.22098499	-0.54758319	2.49355165	1.61537688
## 2023	-0.52070523	1.11406196	-0.08867364	0.21284585
## 2024	-0.45450847	-1.49247939	-0.08907184	-1.73229935
## 2025	-0.60372515	1.12784551	1.11462281	-0.45584863
## 2026	-0.97076347	0.35040089	-1.95474722	-0.63967372
## 2027	-0.40198452	0.10506850	0.04510855	0.84300729
## 2028	0.70200183	0.40813321	0.56816479	0.88684596
## 2029	-0.44914980			

When and why do we need time series models?

When there is *auto-correlation in the residuals* (after modelling trends, seasonality, effects of explanatory variables).

If we ignore autocorrelation:

- ▶ standard error estimates are wrong
- ▶ predictions and prediction intervals are wrong

Terminology / Definitions

time series: y_1, \dots, y_t

autocorrelation: correlation with previous values,

Basic time series processes

AR process

$$\text{AR}(p): \quad x_t = \phi_1 x_{t-1} + \phi_2 x_{t-2} + \dots + \epsilon_t$$

$$\text{AR1:} \quad x_t = \phi x_{t-1} + \epsilon_t$$

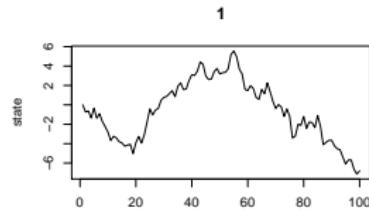
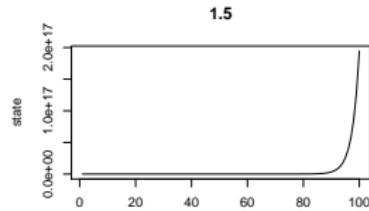
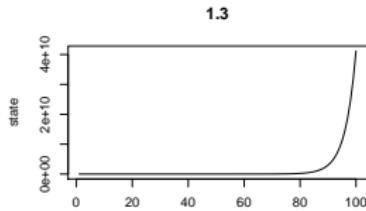
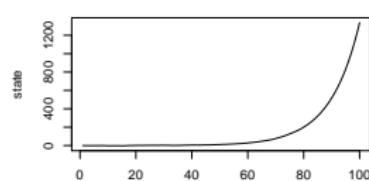
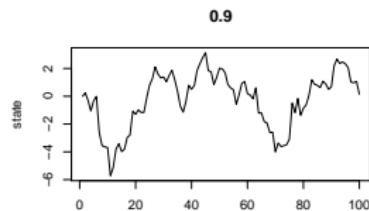
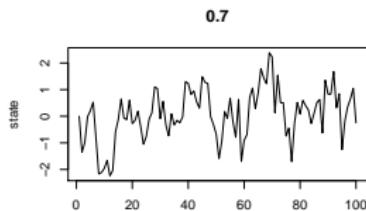
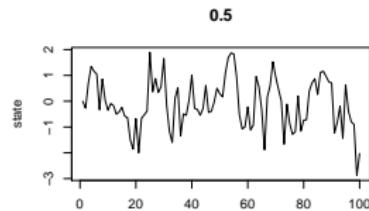
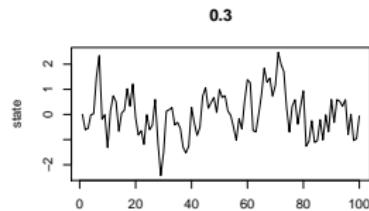
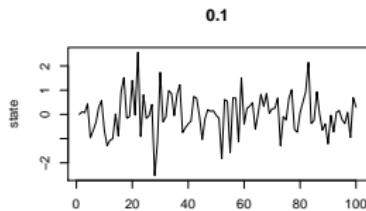
Why would anything behave like this?

I see x_{t-1} as a measure of everything that was not measured explicitly at previous time step.

Random walk

$$x_t = x_{t-1} + \epsilon_t$$

AR1 processes, different ϕ



Stationarity

- ▶ mean, variance, correlations stay constant over time

Why is stationarity important?

There is a single observation per time point. If mean and variance are different for every point, we can't estimate mean and variance, correlation or model parameters.

AR1 processes are stationary if $|\phi| < 1$.

non-stationary means

mean changes, variance changes, seasonality present, correlation changes

MA(q) process

$$y_t = \theta_1 e_{t-1} + \theta_2 e_{t-2} + e_t$$

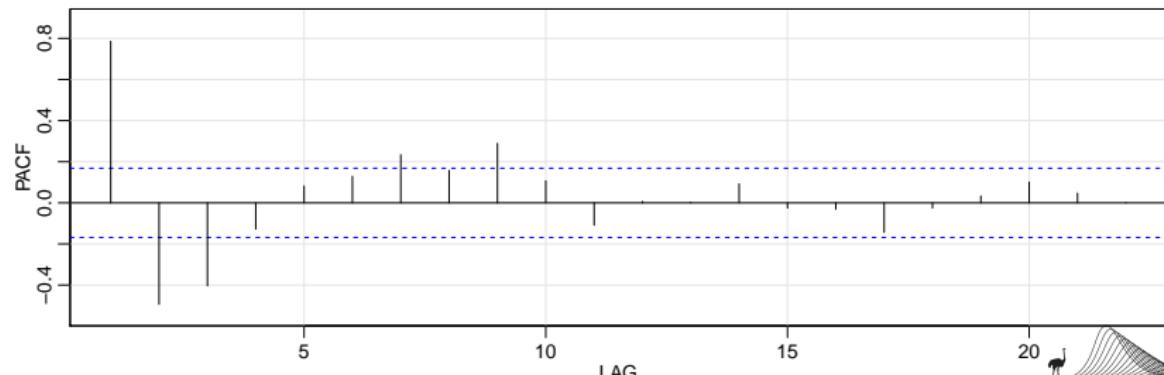
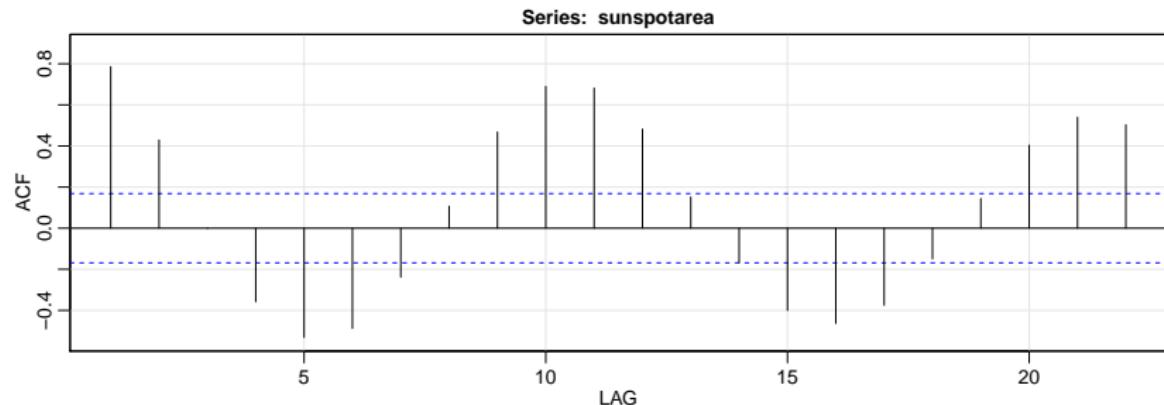
sum of previous shocks / events

White noise

identically, independently distributed, mean 0, no autocorrelation

ACF and PACF

autocorrelation function, partial autocorrelation function



Three approaches to time series modelling

1. ARIMA, very briefly

`arima(p, d, q)`

AR(p), d = difference order, MA(q)

If y_t is not stationary then $y_t - y_{t-1}$ sometimes is (first order differences).

Three approaches to time series modelling

2. Regression

Ignore or model auto-correlation in **errors** like this:

$$y_t = \beta_0 + \beta_1 x_t + \nu_t$$

$$\nu_t = \phi \nu_{t-1} + e_t$$

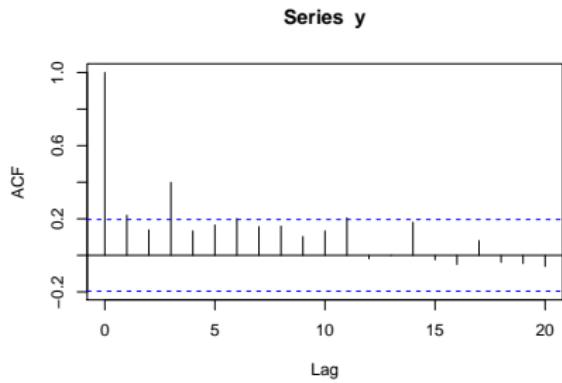
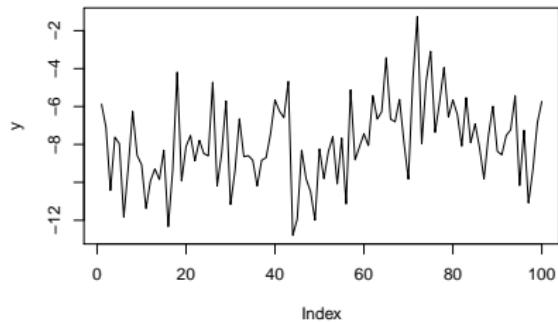
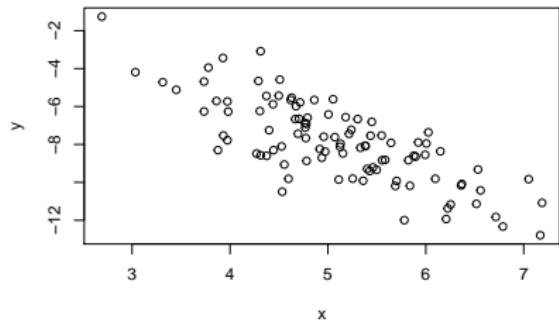
Not this:

$$y_t = \beta_0 + \beta_1 x_t + \phi y_{t-1} + e_t$$

problem: β_1 is not change in response per unit change in x

<https://robjhyndman.com/hyndtsight/arimax/>

Simulate some data with autocorrelation, fit OLS regression model



ARMA errors

arima with xreg models autocorrelation in **errors**

```
a1 <- arima(y, order = c(0, 0, 0), xreg = x)
a2 <- arima(y, order = c(1, 0, 0), xreg = x)
```

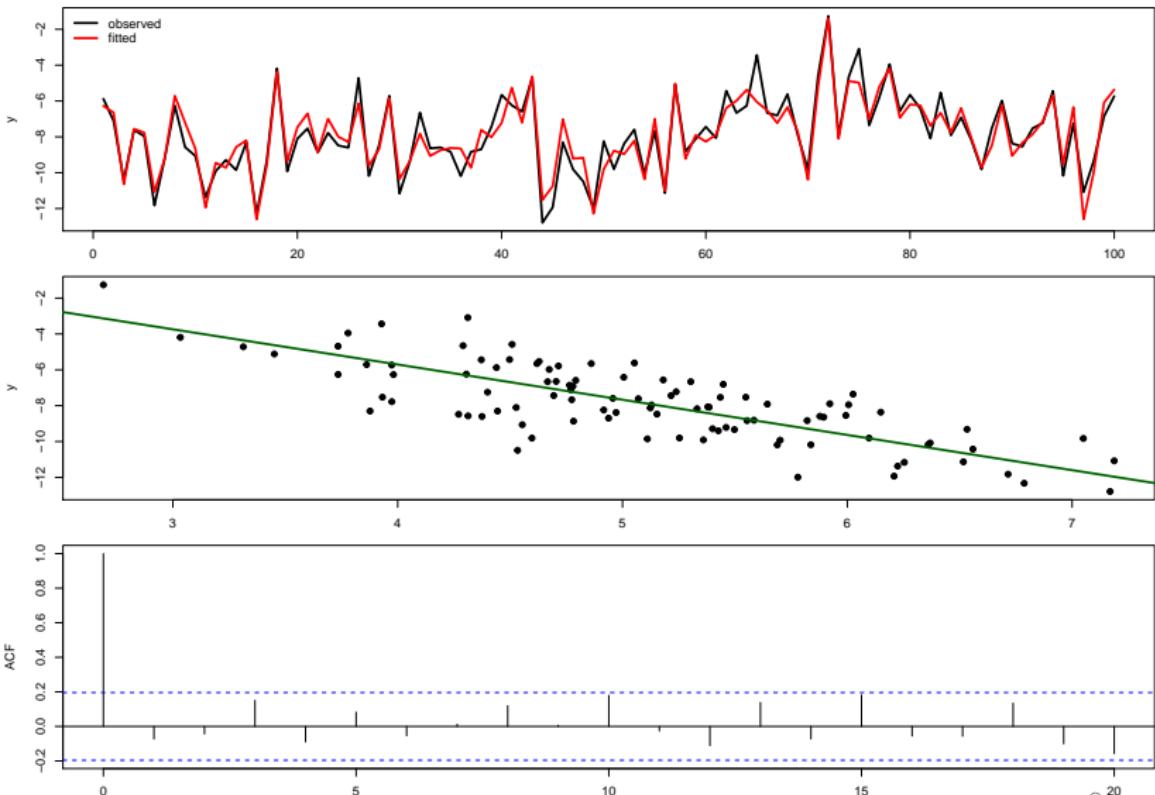
```
##
## Call:
## arima(x = y, order = c(1, 0, 0), xreg = x)
##
## Coefficients:
##             ar1  intercept          x
##             0.8238    2.1524   -1.9644
## s.e.    0.0544    0.5223    0.0640
##
## sigma^2 estimated as 0.5757:  log likelihood = -114.85,
```

true values: $\beta_0 = 3, \beta_1 = -2, \phi = 0.9, \sigma^2 = 0.64$

Regression without ARMA errors

```
##  
## Call:  
## arima(x = y, order = c(0, 0, 0), xreg = x)  
##  
## Coefficients:  
##          intercept           x  
##            1.4424   -1.8305  
## s.e.      0.7726    0.1494  
##  
## sigma^2 estimated as 1.842:  log likelihood = -172.43,
```

ARMA errors



Terminology

fitted values:

$$\hat{y}_t | y_1, y_2, \dots, y_{t-1}$$

usually **one-step-ahead forecast**, using model

residuals:

$$y_t - \hat{y}_t$$

forecasts:

$$y_{t+h} | y_1, y_2, \dots, y_t$$

Point forecasts need **PREDICTION INTERVALS**

Alternative for regression with ARMA errors: GLS

GLS: generalized least squares

linear regression with correlated errors

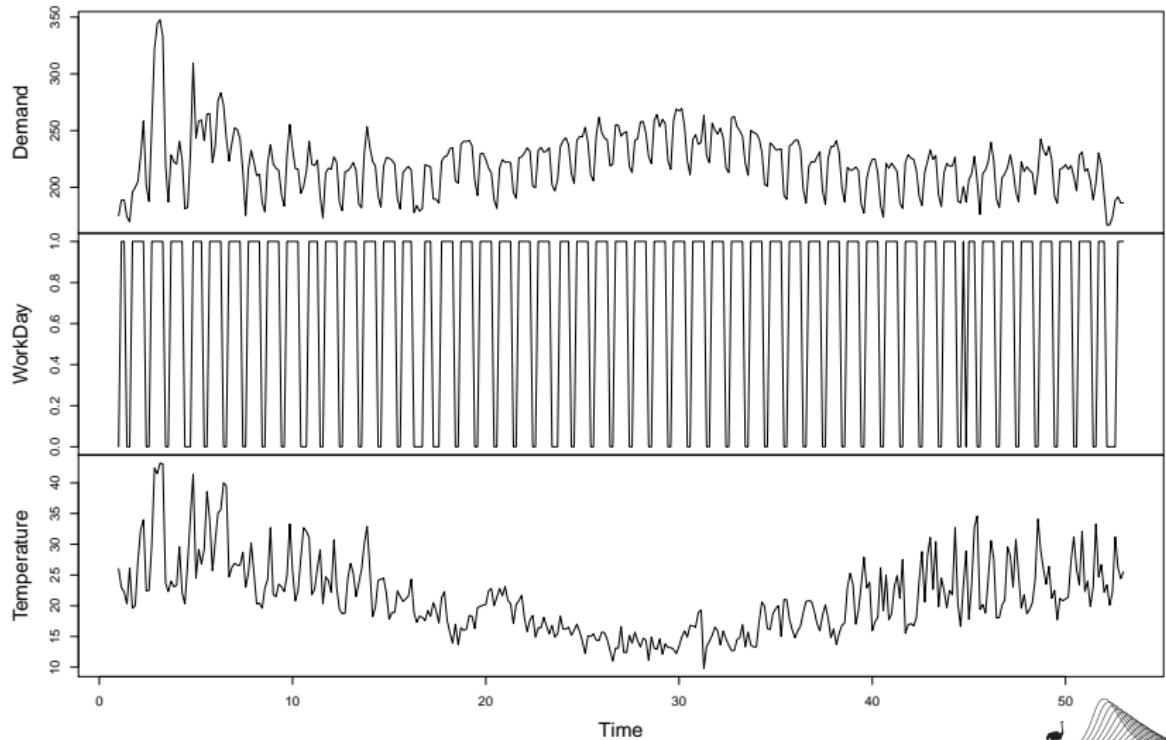
```
library(nlme)
g2 <- gls(y ~ x, correlation = corAR1(form = ~ 1))
```

More flexible alternative for regression with ARMA errors: GAMs with correlated errors

`gamm` (`mgcv`) will fit a GAM, and allow different correlation structures for the errors.

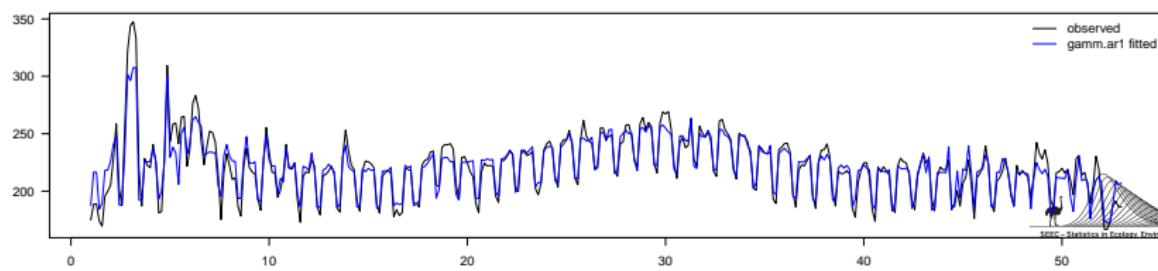
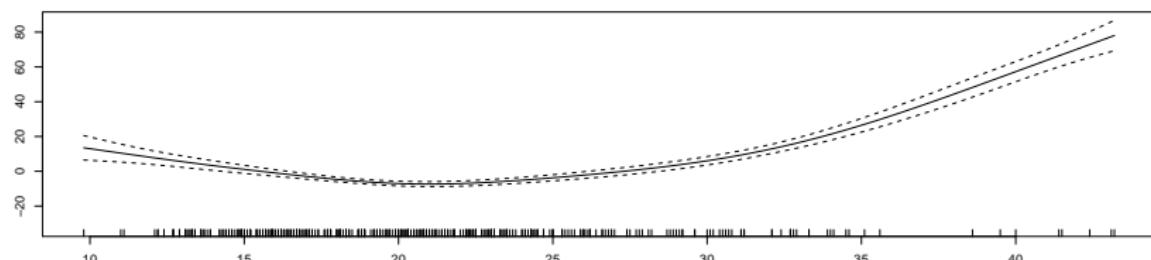
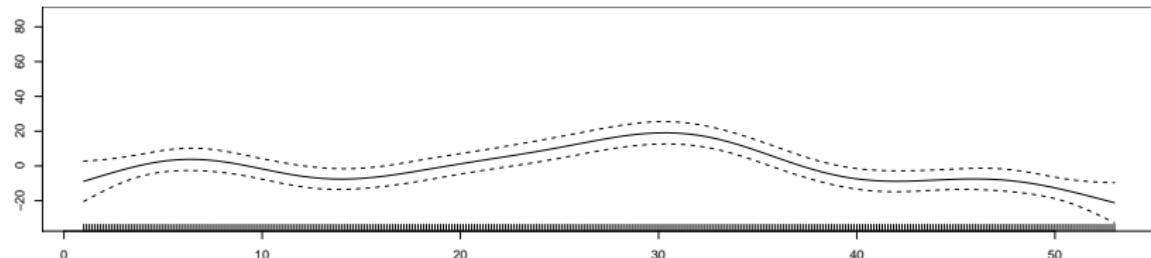
Example: Daily electricity demand (weekday, maximum temperature)

Daily electricity demand, temperature, Victoria, Australia, 2014

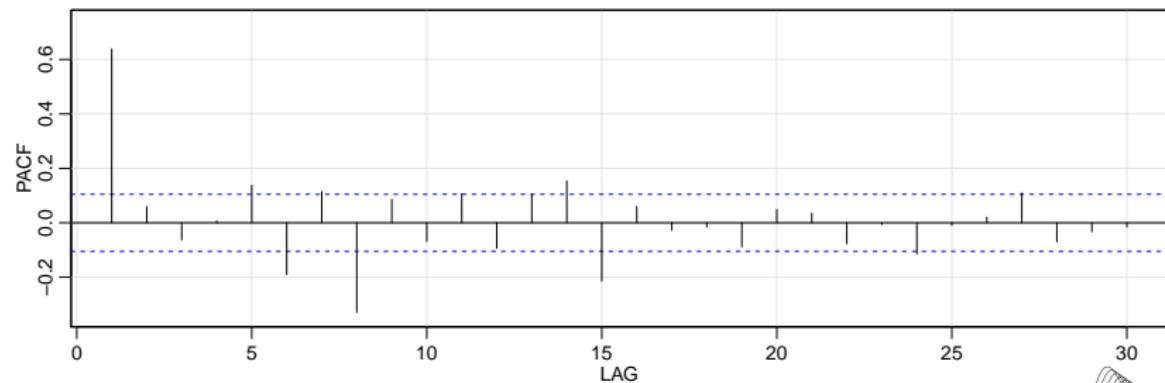
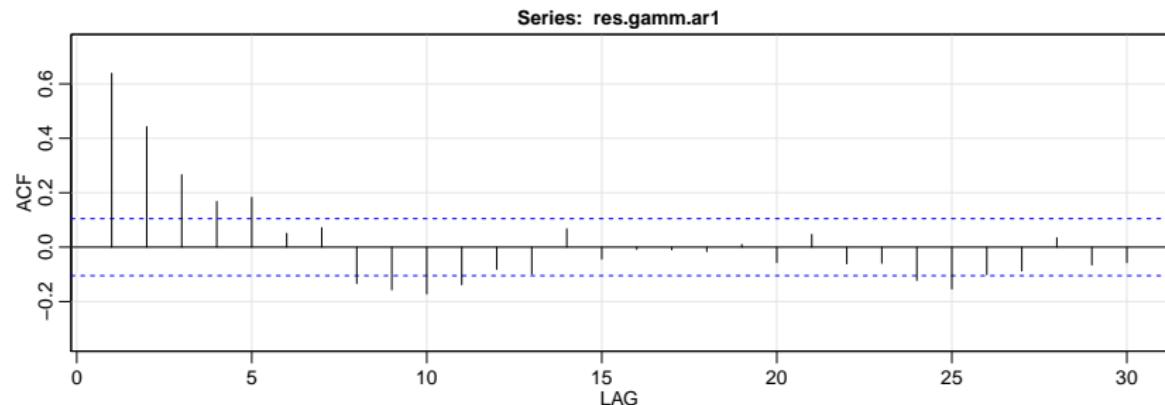


GAMs with correlated errors

```
gamm.ar1 <- gamm(demand ~ s(week) + s(temp) + weekd,  
                   correlation = corAR1(), method = "REML")
```



GAMM residuals



3rd approach: Structural Time Series Models

Trend, season, error modelled explicitly. Similar to regression. Easier to understand.

$$y_t = T_t + S_t + X_t + R_t$$

T_t = trend

S_t = seasonality

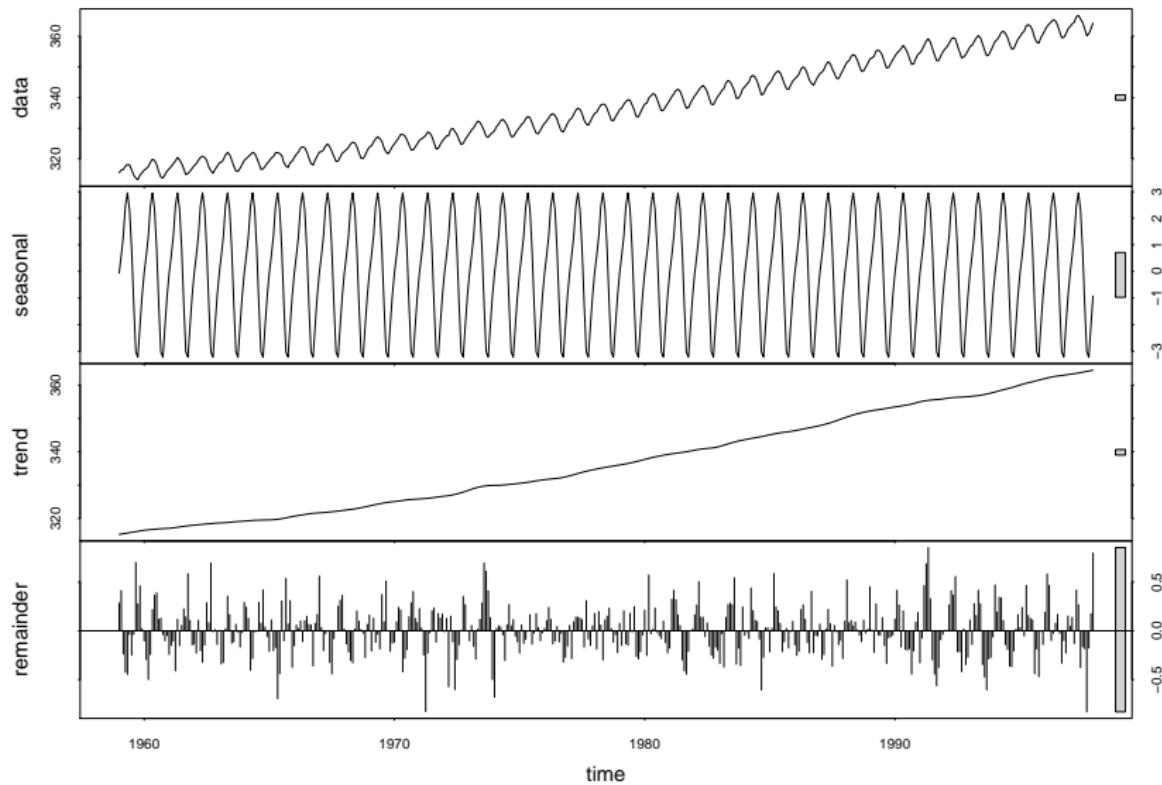
R_t = remainder

X_t = regression terms

International Encyclopedia of Statistical Science:

https://link.springer.com/referenceworkentry/10.1007%2F978-3-642-04898-2_577

decompose



Structural Time Series Models

linear Gaussian state-space models for univariate time series based on a decomposition of the series into components.

Some basic models:

1. Local Level:

$$\mu_t = \mu_{t-1} + \xi_t, \quad \xi_t \sim N(0, \sigma_\xi^2)$$

with observations:

$$y_t = \mu_t + e_t, \quad e_t \sim N(0, \sigma_e^2)$$

trend / mean is a random walk, errors independent

Structural Time Series Models

2. Local Linear Trend:

$$\mu_t = \mu_{t-1} + \beta_t + \xi_t, \quad \xi_t \sim N(0, \sigma_\xi^2)$$

$$\beta_t = \beta_{t-1} + w_t \quad w_t \sim N(0, \sigma_w^2)$$

dynamic / time-varying trend

observations as before

Structural Time Series Models

3. Basic Structural Model

$$y_t = \mu_t + \tau_t + e_t, \quad e_t \sim N(0, \sigma_e^2)$$

τ_t is the seasonal component with dynamics

$$\tau_t = - \sum_{s=1}^{S-1} \tau_{t-s} + w_t, \quad w_t \sim N(0, \sigma_w^2)$$

or

$$\tau_t = \alpha_1 \cos\left(\frac{2\pi t}{\omega}\right) + \alpha_2 \sin\left(\frac{2\pi t}{\omega}\right)$$

State-space model

model the components, can change over time

$$y_t = \underbrace{\mu_t}_{\text{trend}} + \underbrace{\tau_t}_{\text{seasonal}} + \underbrace{X_t \beta_t}_{\text{regression}} + e_t$$

$$\mu_t = \mu_{t-1} + \underbrace{\delta_t}_{\text{slope}} + u_t$$

$$\delta_t = \delta_{t-1} + w_t$$

$$e_t, u_t, w_t$$

independent, (Gaussian) white noise

JAGS

```
## JAGS model for simple state-space model
## local level

model {

## Prior for error of observation process
sigma.obs ~ dunif(0, 10)
tau.obs <- pow(sigma.obs, -2)

## Prior for error of state process
sigma.proc ~ dunif(0, 10)
tau.proc <- pow(sigma.proc, -2)

## State process
mu[1] ~ dnorm(y0, 0.001)

for (t in 2:T) {
  r[t] ~ dnorm(0, tau.proc)
  mu[t] <- mu[t-1] + r[t]
}
```

JAGS

```
## Observation process / model / likelihood
for (t in 1:T) {
  y[t] ~ dnorm(mu[t], tau.obs)
}
```

<https://maialesosky.files.wordpress.com/2016/02/bayesian-state-space-model-applications-for-time-series-analysis-1.pdf>

JAGS: electricity demand, local level state-space model

