

SEEC Toolbox Seminars

Estimating survival from capture-mark-recapture datasets

Res Altwegg

Statistics in Ecology, Environment and Conservation

05 March 2020

Capture-mark-recapture



Non-detection is common



from: Gurney's guide to feathered friends 1969, by E. Gurney and N Gurney

Survival

State process

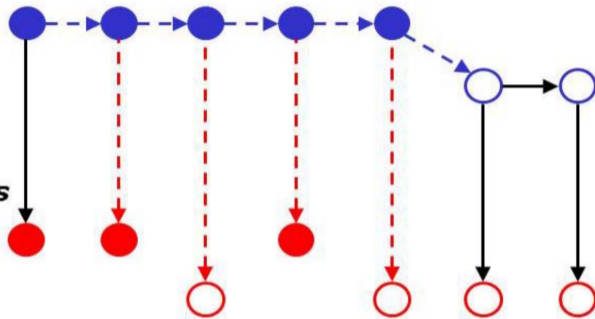
Alive

Dead

Observation process

Seen

Not seen



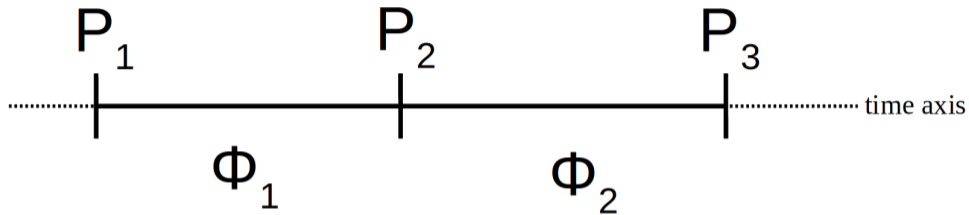
--> Stochastic processes (survival and recapture)

—> Deterministic process

Real world vs model world



CMR experiment

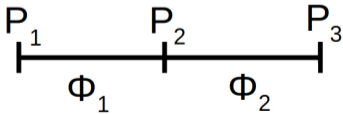


P : recapture probability

Φ : survival probability

Capture histories

Capture history	Meaning	Probability
111	ringed at occasion 1, bird survived first period, seen at occasion 2, bird survived second period, seen at occasion 3	$\Phi_1 P_2 \Phi_2 P_3$
101	ringed at occasion 1, bird survived first period, not seen at occasion 2, bird survived second period, seen at occasion 3	$\Phi_1 (1 - P_2) \Phi_2 P_3$
110	ringed at occasion 1, bird survived first period, seen at occasion 2, bird survived second period, not seen at occasion 3 or bird did not survive second period	$\Phi_1 P_2 [\Phi_2 (1 - P_3) + (1 - \Phi_2)]$ $= \Phi_1 P_2 (1 - \Phi_2 P_3)$

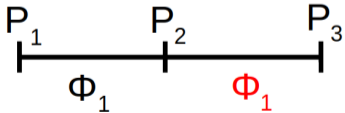


Parameters P and Φ estimated
by maximum likelihood

P_1 not estimated!

Capture histories

Capture history	Meaning	Probability
111	ringed at occasion 1, bird survived first period, seen at occasion 2, bird survived second period, seen at occasion 3	$\Phi_1 P_2 \Phi_1 P_3$
101	ringed at occasion 1, bird survived first period, not seen at occasion 2, bird survived second period, seen at occasion 3	$\Phi_1 (1 - P_2) \Phi_1 P_3$
110	ringed at occasion 1, bird survived first period, seen at occasion 2, bird survived second period, not seen at occasion 3 or bird did not survive second period	$\Phi_1 P_2 [\Phi_1 (1 - P_3) + (1 - \Phi_1)]$ $= \Phi_1 P_2 (1 - \Phi_1 P_3)$



Parameters P and Φ estimated by maximum likelihood

P_1 not estimated!

Program MARK

A Gentle Introduction

EVAN G. COOCH & GARY C. WHITE (eds.)

◆ 15th edition

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08/14/2018

Data wrangling: hadeda data

```
head(ch) # here we've got our capture history
```

```
##  
##      2006 2007 2008 2009 2010 2011 2012 2013 2014 2015 2016 2017  
## -      0    0    1    0    0    0    0    0    0    0    0    0  
## AA     1    0    0    0    0    0    0    0    0    0    0    0  
## AB     1    0    0    0    0    0    0    0    0    0    0    0  
## AC     1    0    0    0    0    0    0    0    0    0    0    0  
## AD     1    0    0    0    0    0    0    0    0    0    0    0  
## AF     1    0    0    0    0    0    0    0    0    0    0    0
```

```
dim(ch) # we have data on 261 individuals for 12 years
```

```
## [1] 261 12
```

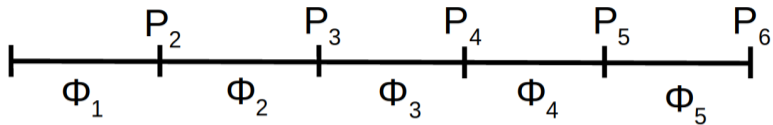


Input format for MARK

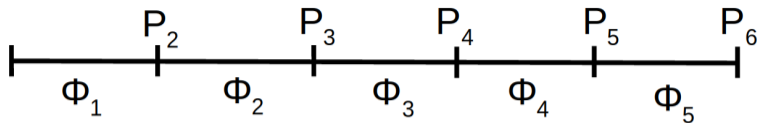
```
write.table(cbind(ch," ",1,";"), "hadedada.inp", sep="",  
            row.names=F, col.names=F, quote=F)
```

```
00100000000000 1;  
10000000000000 1;  
10000000000000 1;  
10000000000000 1;  
10000000000000 1;  
10000000000000 1;
```

Parameter Index Matrix



Parameter Index Matrix



1	2	3	4	5
	2	3	4	5
		3	4	5
			4	5
				5

- cohort 1
- cohort 2
- cohort 3
- cohort 4
- cohort 5

“PIM”

MARK analysis using RMark

```
library(RMark)

# MARK input file with capture histories
ch1 <- apply(ch, 1 , paste , collapse = "" )
hade <- data.frame(ch=ch1, group = 1)
hade$ch <- as.character(hade$ch)

# run default model Phi()P()
model1 <- mark(hade)
```

```
##
## Output summary for CJS model
## Name : Phi(~1)p(~1)
##
## Npar : 2
## -2lnL: 964.9276
## AICc : 968.9553
```

Output

Beta

	estimate	se	lcl	ucl
Phi:(Intercept)	0.5803299	0.1004607	0.3834269	0.7772329
p:(Intercept)	-0.3789708	0.1363474	-0.6462117	-0.1117299

Real Parameter Phi

	1	2	3	4	5	6	7	
1	0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0
2		0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0
3			0.6411433	0.6411433	0.6411433	0.6411433	0.6411433	0
4				0.4063752	0.4063752	0.4063752	0.4063752	0
5					0.4063752	0.4063752	0.4063752	0

Take a bit more control

```
hade.process <- process.data(hade,model="CJS", begin.time=2006)
hade.ddl <- make.design.data(hade.process)
head(hade.ddl$Phi)
```

##	par.index	model.index	group	cohort	age	time	occ.cohort	Cohort	Age	Time
## 1	1	1	1	2006	0	2006	1	0	0	0
## 2	2	2	1	2006	1	2007	1	0	1	1
## 3	3	3	1	2006	2	2008	1	0	2	2
## 4	4	4	1	2006	3	2009	1	0	3	3
## 5	5	5	1	2006	4	2010	1	0	4	4
## 6	6	6	1	2006	5	2011	1	0	5	5

Working with constraints

“all different” PIM

1	2	3	4	5
	6	7	8	9
		10	11	12
			13	14
				15

$$\text{logit}(\Phi) = \beta_0 + \sum \beta_i \times x_i$$

x_i are covariates

- ▶ time, age, sex
- ▶ environmental, e.g. rain
- ▶ individual, e.g. size at fledging

β are parameters to be estimated

“Design Data”

- ▶ Mapping x_i onto “all different” PIM
- ▶ Parameter number in rows
- ▶ x in columns

```
head(hade.ddl$Phi)
```

```
##      par.index model.index group cohort age time occ.cohort Cohort Age Time
## 1           1           1      1   2006  0 2006           1      0  0
## 2           2           2      1   2006  1 2007           1      0  1
## 3           3           3      1   2006  2 2008           1      0  2
## 4           4           4      1   2006  3 2009           1      0  3
## 5           5           5      1   2006  4 2010           1      0  4
## 6           6           6      1   2006  5 2011           1      0  5
```



Prepare some model structures and run the fully time-dependent model

```
dot <- list(formula=~1)
time <- list(formula=~time)
age <- list(formula=~age)
timeage <- list(formula=~time + age)

# runs the fully time-dependent model
Phi.time.P.time <- mark(hade.process,hade.ddl,
                        model.parameters=list(Phi=time, p=time))
```

Summary output

```
summary(Phi.time.P.time)
```

Output summary for CJS model

Name : Phi(~time)p(~time)

Npar : 22 (unadjusted=21)

-2lnL: 930.9508

AICc : 977.4012 (unadjusted=975.18273)

Beta

	estimate	se	lcl	ucl
Phi:(Intercept)	1.0822175	1.3642741	-1.5917597	3.7561947

...

Have a look at the parameter estimates

```
Phi.time.P.time$results$real # parameter estimates
```

					estimate	se	lcl	ucl
Phi	g1	c2006	a0	t2006	0.7469134	0.2578939	0.1691365	0.9771613
Phi	g1	c2006	a1	t2007	0.3558946	0.0828915	0.2138338	0.5288463
Phi	g1	c2006	a2	t2008	0.8116639	0.1377452	0.4242693	0.9618373
...								
Phi	g1	c2006	a10	t2016	0.7236854	0.0000000	0.7236854	0.7236854
p	g1	c2006	a1	t2007	0.3966944	0.1692755	0.1411761	0.7245293
p	g1	c2006	a2	t2008	0.3959044	0.1101140	0.2099931	0.6177118
...								
p	g1	c2006	a11	t2017	0.5922065	0.0000000	0.5922065	0.5922065



Counting parameters. . .

- ▶ MARK tries to work out how many parameters were estimable
- ▶ RMark assumes all specified parameters are estimable
- ▶ You need to check and adjust if necessary

```
Phi.time.P.time <- adjust.parameter.count(Phi.time.P.time, 21)
```

```
##  
## Number of parameters adjusted from 22 to 21  
  
## Adjusted AICc = 975.182724057971  
## Unadjusted AICc = 975.18273
```

Run some more models

```
Phi.dot.P.dot <- mark(hade.process,hade.ddl,  
                      model.parameters=list(Phi=dot, p=dot))  
Phi.age.P.dot <- mark(hade.process,hade.ddl,  
                      model.parameters=list(Phi=age, p=dot))  
Phi.age.P.time <- mark(hade.process,hade.ddl,  
                       model.parameters=list(Phi=age, p=time))  
Phi.timeage.P.dot <- mark(hade.process,hade.ddl,  
                          model.parameters=list(Phi=timeage, p=dot))  
Phi.timeage.P.time <- mark(hade.process,hade.ddl,  
                           model.parameters=list(Phi=timeage, p=time))
```

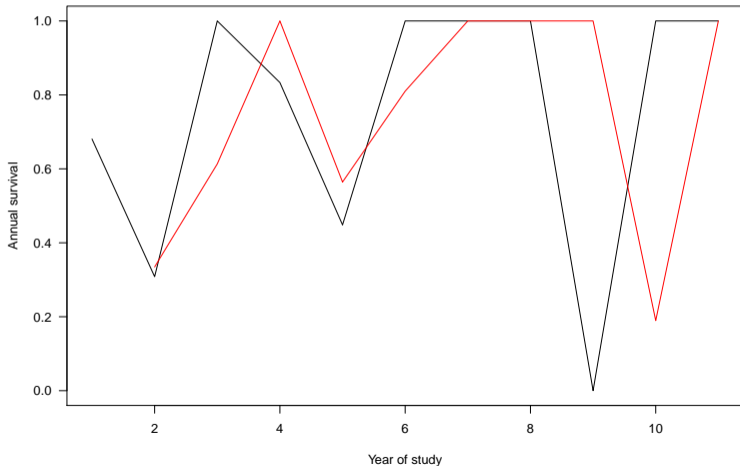
Model selection

```
hade.cjs.results <- collect.models()  
hade.cjs.results
```

##	model	npar	AICc	DeltaAICc	weight	Deviance
## 5	Phi(~time + age)p(~1)	22	950.0441	0.00000	9.955527e-01	250.3671
## 1	Phi(~age)p(~1)	12	961.5801	11.53601	3.112092e-03	283.6159
## 6	Phi(~time + age)p(~time)	32	963.4889	13.44478	1.198304e-03	241.0215
## 3	Phi(~1)p(~1)	2	968.9553	18.91115	7.790406e-05	311.7009
## 2	Phi(~age)p(~time)	22	969.6318	19.58769	5.554582e-05	269.9548
## 4	Phi(~time)p(~time)	21	975.1827	25.13858	3.461711e-06	277.7241

Plotting results

```
plot(1:11, Phi.timeage.P.dot$results$real$estimate[1:11], type='l', las=1,  
     ylab = "Annual survival", xlab = "Year of study")  
lines(2:11, Phi.timeage.P.dot$results$real$estimate[12:21], col='red')
```



Revisiting age

```
head(hade.ddl$Phi)
```

##	par.index	model.index	group	cohort	age	time	occ.cohort	Cohort	Age	Time
## 1	1	1	1	2006	0	2006	1	0	0	0
## 2	2	2	1	2006	1	2007	1	0	1	1
## 3	3	3	1	2006	2	2008	1	0	2	2
## 4	4	4	1	2006	3	2009	1	0	3	3
## 5	5	5	1	2006	4	2010	1	0	4	4
## 6	6	6	1	2006	5	2011	1	0	5	5

```
ageclass <- NA  
ageclass[hade.ddl$Phi$Age == 0] <- "juvenile"  
ageclass[hade.ddl$Phi$Age == 1] <- "immature"  
ageclass[hade.ddl$Phi$Age > 1] <- "adult"  
hade.ddl$Phi$ageclass <- as.factor(ageclass)
```



Look at the design data

```
head(hade.ddl$Phi, n=15)
```

	par.index	model.index	group	cohort	age	time	occ.cohort	Cohort	Age	Time	ageclass
1	1	1	1	2006	0	2006	1	0	0	0	juvenile
2	2	2	1	2006	1	2007	1	0	1	1	immature
3	3	3	1	2006	2	2008	1	0	2	2	adult
4	4	4	1	2006	3	2009	1	0	3	3	adult
5	5	5	1	2006	4	2010	1	0	4	4	adult
6	6	6	1	2006	5	2011	1	0	5	5	adult
7	7	7	1	2006	6	2012	1	0	6	6	adult
8	8	8	1	2006	7	2013	1	0	7	7	adult
9	9	9	1	2006	8	2014	1	0	8	8	adult
10	10	10	1	2006	9	2015	1	0	9	9	adult
11	11	11	1	2006	10	2016	1	0	10	10	adult
12	12	12	1	2007	0	2007	2	1	0	1	juvenile
13	13	13	1	2007	1	2008	2	1	1	2	immature
14	14	14	1	2007	2	2009	2	1	2	3	adult
15	15	15	1	2007	3	2010	2	1	3	4	adult

Run models with age class

```
ageclass <- list(formula=~ageclass)
timeageclass <- list(formula=~time + ageclass)

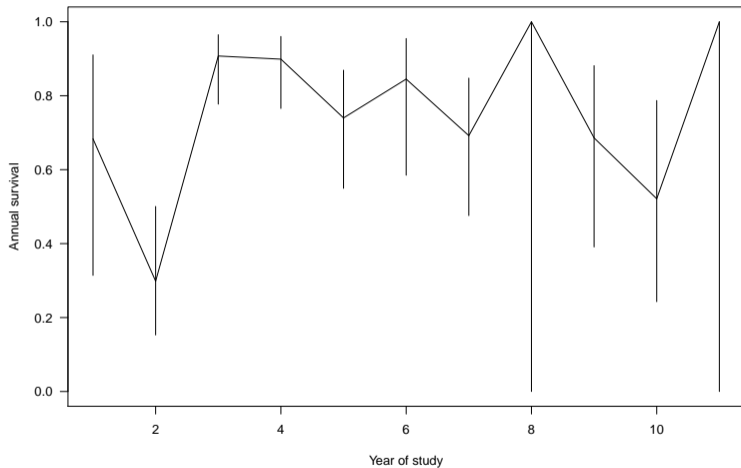
Phi.ageclass.P.dot <- mark(hade.process,hade.ddl,
                           model.parameters=list(Phi=ageclass, p=dot))
Phi.timeageclass.P.dot <- mark(hade.process,hade.ddl,
                               model.parameters=list(Phi=timeageclass, p=dot))
```

Model selection

```
hade.cjs.results <- collect.models()
hade.cjs.results
```

##	model	npar	AICc	DeltaAICc	weight	Deviance
## 8	Phi(~time + ageclass)p(~1)	14	947.8054	0.000000	5.936613e-01	265.5811
## 3	Phi(~ageclass)p(~1)	4	949.8681	2.062723	2.116529e-01	288.5486
## 6	Phi(~time + age)p(~1)	22	950.0441	2.238758	1.938200e-01	250.3671
## 1	Phi(~age)p(~1)	12	961.5801	13.774764	6.058802e-04	283.6159
## 7	Phi(~time + age)p(~time)	32	963.4889	15.683540	2.332928e-04	241.0215
## 4	Phi(~1)p(~1)	2	968.9553	21.149909	1.516681e-05	311.7009
## 2	Phi(~age)p(~time)	22	969.6318	21.826448	1.081398e-05	269.9548
## 5	Phi(~time)p(~time)	21	975.1827	27.377339	6.739459e-07	277.7241

Plotting results



Assumptions

1. every marked animal present in the population at time (i) has the same probability of recapture (p_i)
2. every marked animal in the population immediately after time (i) has the same probability of surviving to time ($i+1$)
3. marks are not lost or missed.
4. all samples are instantaneous, relative to the interval between occasion (i) and ($i+1$), and each release is made immediately after the sample.

Goodness-of-fit test: 'Test 3'

- ▶ Assumption: all marked individuals have same survival probability.
- ▶ 'Of those individuals seen at occasion i , how many were ever seen again?'

seen before (i)	seen again	not seen again
no	f	f
yes	f	f

GOF testing

```
require(R2ucare) # load U-CARE
```

```
ch.gof <- as.matrix(ch)
```

```
freq <- rep(1, length=dim(ch.gof)[1])
```

```
overall_CJS(ch.gof,freq)
```

```
##                ##          chi2 degree_of_freedom p_value  
## Gof test for CJS model: 83.231                31          0
```

Detailed tests

```
test3sr(ch.gof,freq)
```

```
$test3sr
  stat      df    p_val sign_test
17.894 10.000  0.057   2.513
```

```
$details
 component stat p_val signed_test test_perf
1          2 0.182 0.67   -0.427   Fisher
2          3 4.027 0.045    2.007 Chi-square
...
```

```
test3sm(ch.gof,freq)
```

```
$test3sm
  stat  df p_val
5.614 6.000 0.468
```

```
$details
 component stat df p_val test_perf
1          2 0.936 1 0.333   Fisher
2          3 3.155 1 0.076 Chi-square
...
```

```
test2ct(ch.gof,freq)
```

```
$test2ct
  stat      df    p_val sign_test
53.996  9.000  0.000   -6.705
```

```
$details
 component dof  stat p_val signed_test test_perf
1          2  1  3.452 0.063   -1.858   Fisher
2          3  1  8.155 0.004   -2.856 Chi-square
...
```

```
test2cl(ch.gof,freq)
```

```
$test2cl
  stat  df p_val
5.727 6.000 0.454
```

```
$details
 component dof  stat p_val test_perf
1          2  1  1.913 0.167   Fisher
2          3  1  0.748 0.387 Chi-square
...
```

M-array

```
marray(ch.gof, freq)
```

\$R		\$m													\$never
	[,1]		[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]	[,8]	[,9]	[,10]	[,11]		[,1]
[1,]	27	[1,]	8	0	1	2	0	0	0	0	0	0	0	[1,]	16
[2,]	73	[2,]	0	12	0	2	1	0	1	0	1	0	1	[2,]	55
[3,]	59	[3,]	0	0	19	3	3	2	0	1	1	0	0	[3,]	30
[4,]	95	[4,]	0	0	0	32	5	0	0	2	1	0	0	[4,]	55
[5,]	65	[5,]	0	0	0	0	17	2	1	3	0	0	0	[5,]	42
[6,]	29	[6,]	0	0	0	0	0	15	1	1	0	0	0	[6,]	12
[7,]	29	[7,]	0	0	0	0	0	0	9	1	3	0	0	[7,]	16
[8,]	16	[8,]	0	0	0	0	0	0	0	11	0	0	0	[8,]	5
[9,]	20	[9,]	0	0	0	0	0	0	0	0	8	0	1	[9,]	11
[10,]	16	[10,]	0	0	0	0	0	0	0	0	0	6	1	[10,]	9
[11,]	7	[11,]	0	0	0	0	0	0	0	0	0	0	3	[11,]	4



Measure of overdispersion

```
chi2 <- overall_CJS(ch.gof,freq)$chi2 # test statistic  
df <- overall_CJS(ch.gof,freq)$degree_of_freedom # degrees of freedom  
chi2 / df
```

```
## [1] 2.684871
```

Reasons for lack of fit and possible solutions

- ▶ Test 2 significant:
 - ▶ 'trap response' → 'time since capture' models
 - ▶ non-random temporary emigration → multi-state models
- ▶ Test 3 significant:
 - ▶ age effect → age models
 - ▶ cohorts differ → cohort models
 - ▶ transients → 'time since first capture' models
- ▶ Non-systematic deviations, general overdispersion → apply overdispersion factor, \hat{C} if below ~ 3



SEEC STATS TOOLBOX 2020 SCHEDULE

Date	Topic	Speaker
13 February	Spatio-temporal point pattern analysis	Mzabalazo Ngwenya
5 March	Capture-recapture	Res Altwegg
16 April	Reproducible research	Glenn Moncrieff
7 May	TBD	TBD
18 June	TBD	TBD
9 July	TBD	TBD
30 July	Time-to-first detection (TTD) occupancy models	Dominic Henry
20 August	TBD	TBD
10 September	MCMC	Allan Clark
1 October	Spatial GLMMs	David Maphisa
12 November	Species richness and diversity	Jasper Slingsby