

Multimodel inference (model selection and model averaging)

SEEC Toolbox Seminar

Res Altwegg

27/05/2021

Overview

- ▶ Ranking a set of model
- ▶ Multiple working hypotheses
- ▶ Model averaging



<http://www.seec.uct.ac.za/>

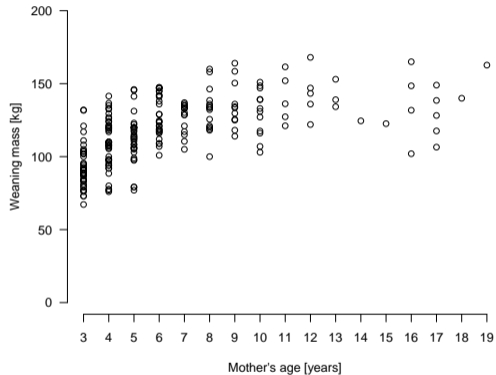
@SEEC_UCT

Good statistical models

Should fit the structure in the data but not the noise.

- ▶ Underfitting: failure to fit structure in the data → prediction bias
- ▶ Overfitting: fits to noise → loss of precision

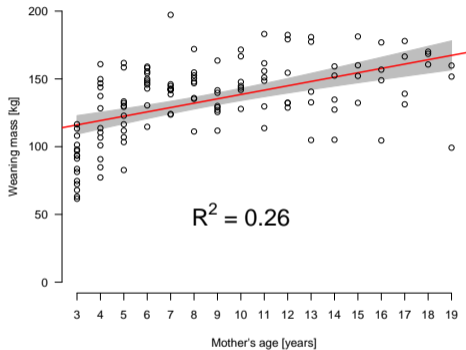
Elephant seals on Marion Island



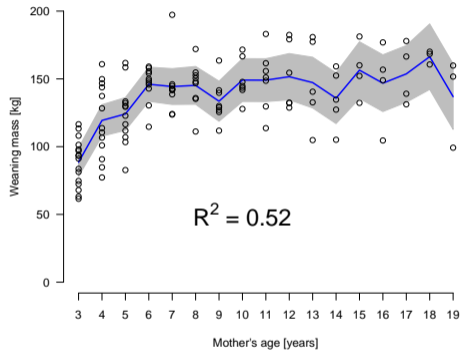
Thanks to Chris Oosthuizen and the Marion Island Marine Mammal Program, University of Pretoria (www.marionseals.com)



Elephant seals

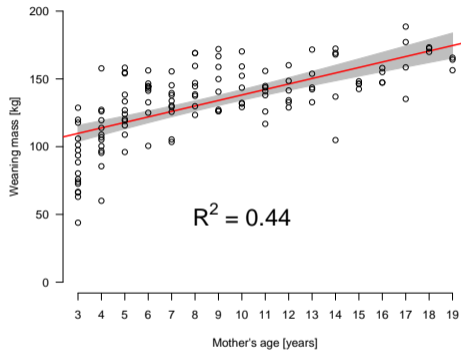


$$Y_j = \beta_0 + \beta_1 \times X_j + \epsilon_j$$

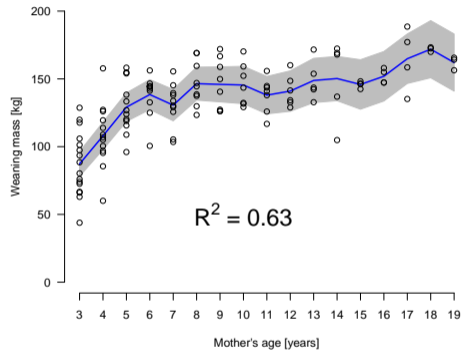


$$Y_{ij} = \mu_j + \epsilon_{ij}$$

Elephant seals

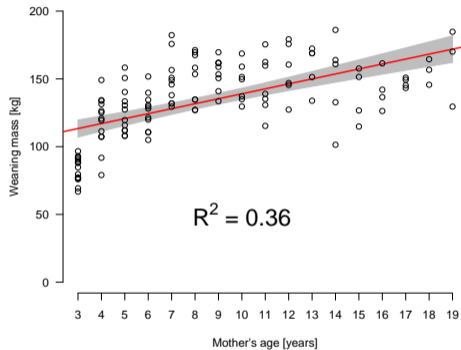


$$Y_j = \beta_0 + \beta_1 \times X_j + \epsilon_j$$

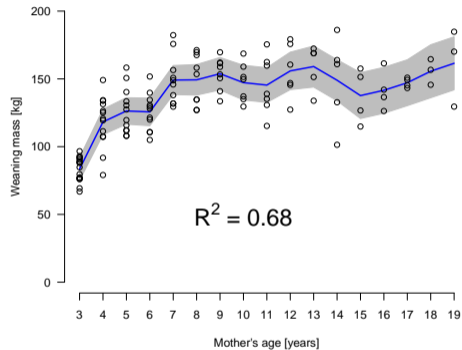


$$Y_{ij} = \mu_j + \epsilon_{ij}$$

Elephant seals

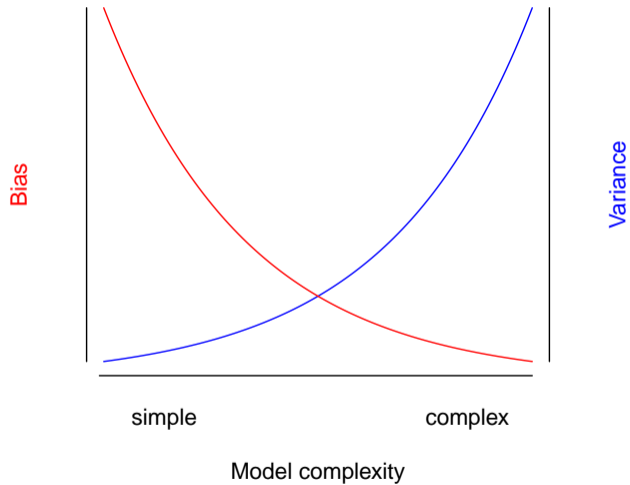


$$Y_j = \beta_0 + \beta_1 \times X_j + \epsilon_j$$



$$Y_{ij} = \mu_j + \epsilon_{ij}$$

Bias-variance trade-off

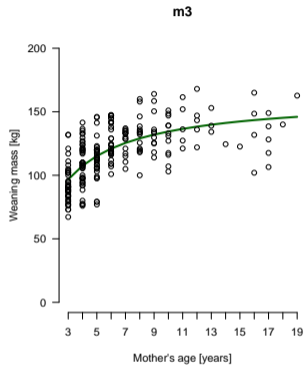
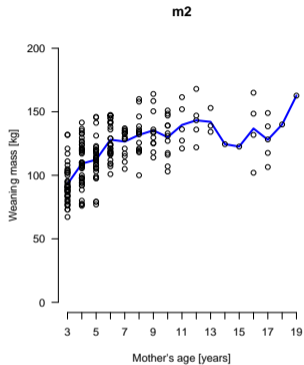
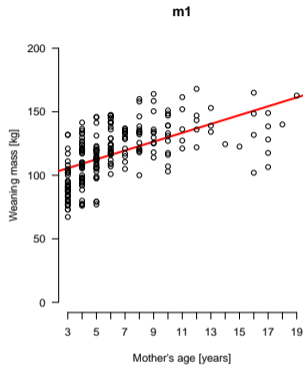


Akaike's Information Criterion

$$AIC = -2 \log(\mathcal{L}(\hat{\theta}|\text{Data})) + 2K$$

- ▶ Best balance between bias and variance
- ▶ Smaller value is better

Model selection analysis

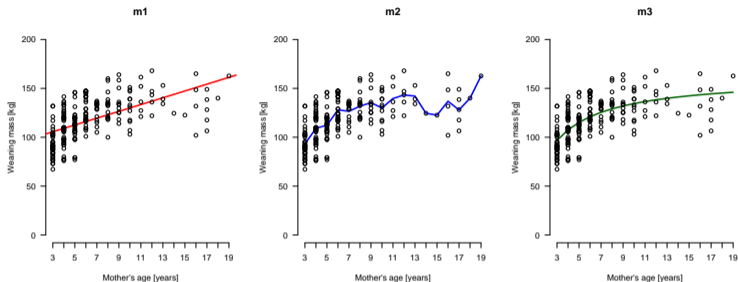


$$Y_i = \beta_0 + \beta_1 \times X_i + \epsilon_i$$

$$Y_{ij} = \mu_j + \epsilon_{ij}$$

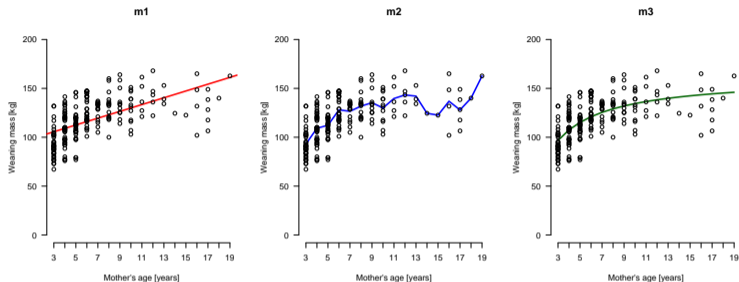
$$Y_i = \frac{\beta_0 \times X_i}{1 + \beta_0 \times \beta_1 \times X_i} + \epsilon_i$$

Model selection analysis



	$-2 \times \log \text{lik}$	K	AIC
m1	1940.50	3	1946.50
m2	1868.14	18	1904.14
m3	1891.49	3	1897.49

Model selection analysis



	$-2 \times \log\text{lik}$	K	AIC	Δ AIC
m1	1940.50	3	1946.50	49.00
m2	1868.14	18	1904.14	6.65
m3	1891.49	3	1897.49	0

Akaike weights

The weight of evidence in favor of model i being the best in the set:

$$w_i = \frac{\exp(-\frac{1}{2}\Delta_i)}{\sum_{r=1}^R \exp(-\frac{1}{2}\Delta_r)}$$

	$-2 \times \text{loglik}$	K	AIC	Δ AIC	w
m1	1940.50	3	1946.50	49.00	0.00
m2	1868.14	18	1904.14	6.65	0.03
m3	1891.49	3	1897.49	0	0.97

Model m3 had 97% of the support relative to the other models.

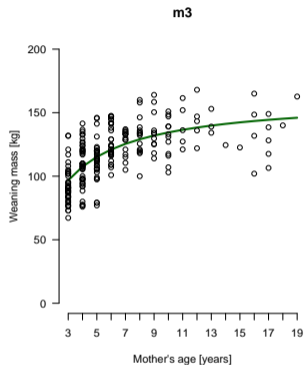
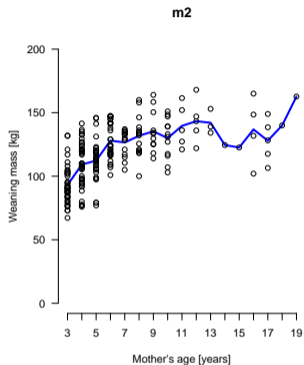
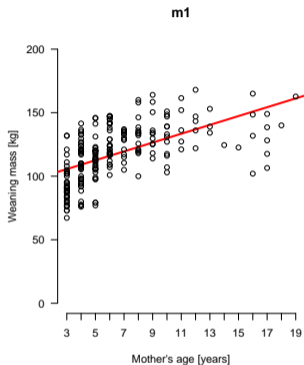
Evidence ratios

$$\text{Evidence ratio} = \frac{w_i}{w_j}$$

	$-2 \times \text{loglik}$	K	AIC	Δ AIC	w
m1	1940.50	3	1946.50	49.00	0.00
m2	1868.14	18	1904.14	6.65	0.03
m3	1891.49	3	1897.49	0	0.97

$$\frac{w_3}{w_2} = \frac{0.97}{0.03} = 27.8$$

Model m3 was 28 times more likely than m2 to be the best in the set.



$$Y_i = \beta_0 + \beta_1 \times X_i + \epsilon_i$$

$$Y_{ij} = \mu_j + \epsilon_{ij}$$

$$Y_i = \frac{\beta_0 \times X_i}{1 + \beta_0 \times \beta_1 \times X_i} + \epsilon_i$$

Model selection analysis

Two different scientific goals need fundamentally different approaches.

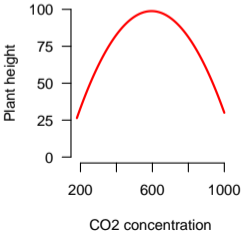
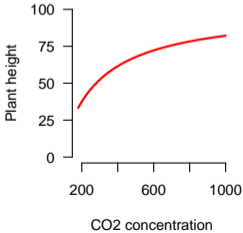
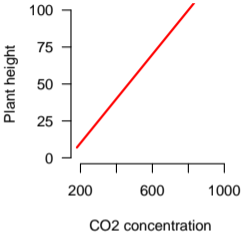
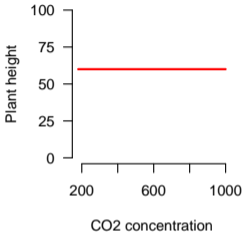
- ▶ Hypothesis-based research
- ▶ Data mining / hypothesis generation

Statistical Modelling / Scientific Research: The steps

1. come up with a set of biological hypotheses
2. translate the hypotheses into statistical models
3. data collection: field, experiment, observations
4. fit the models to these data
5. evaluate the relative support each model (hypothesis) gets from the data
6. answer the biological question

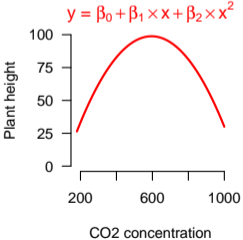
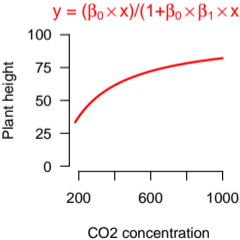
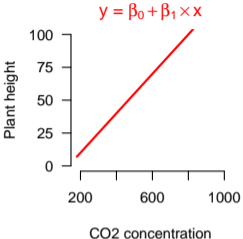
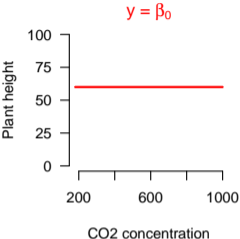
Step 1: Formulate biological hypotheses

How does [CO₂] affect growth of *Acacia karroo*?



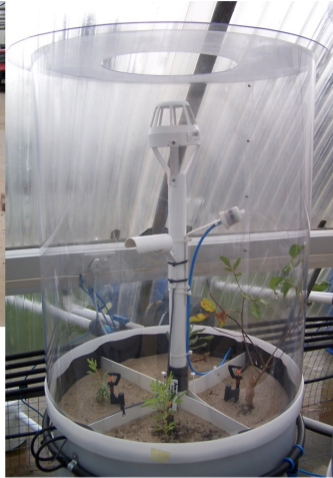
"Acacia karroo, bloeityd, Roodeplaat NR" by JMK - Own work. Licensed under CC BY-SA 3.0 via Wikimedia Commons

Step 2: Translate hypotheses into models



"Acacia karroo, bloeityd, Roodeplaat NR" by JMK - Own work. Licensed under CC BY-SA 3.0 via Wikimedia Commons

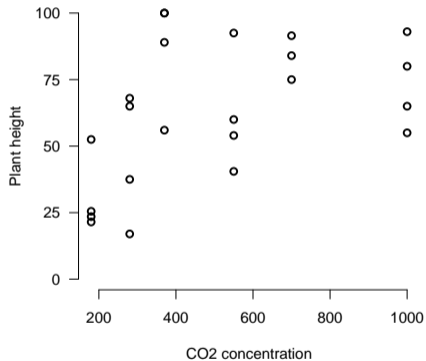
Step 3: Data collection



Kgope et al 2009. *Austral Ecology* 35:451–463.

Fotos: Guy F Midgley

Step 3: Data collection

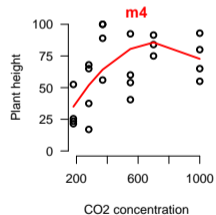
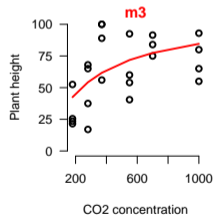
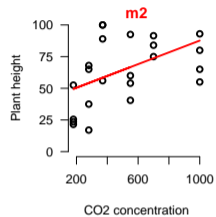
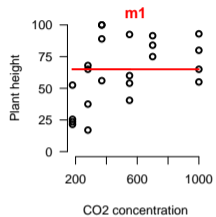


Kgope et al 2009. *Austral Ecology* 35:451–463.

Step 4: Fit models to data

```
m1<-lm(Total.stem.length~1)
m2<-lm(Total.stem.length~CO2)
m3<-nls(Total.stem.length~(b0*CO2)/(1+b0*b1*CO2),
        start=list(b0=1,b1=0.01))
m4<-lm(Total.stem.length~CO2+I(CO2^2))
```

Step 4: Fit models to data



Step 5: Evaluate relative support

```
aics <- AIC(m1,m2,m3,m4)
delta.aics <- aics$AIC - min(aics$AIC)
wi <- exp(-0.5*delta.aics)/sum(exp(-0.5*delta.aics))

logliks <- c(logLik(m1),logLik(m2),logLik(m3),logLik(m4))

models <- c("m1","m2","m3","m4")
mstable <- data.frame(models, -2*logliks, aics$df,
                      aics$AIC, delta.aics,wi)
```


Step 5: Evaluate relative support

	$-2 \times \log\text{lik}$	K	AIC	Δ AIC	w
m1	226.56	2	230.56	8.58	0.01
m2	220.44	3	226.45	4.46	0.06
m3	216.60	3	222.60	0.61	0.40
m4	213.99	4	221.99	0.00	0.54

Step 6: Answer biological question

	$-2 \times \log\text{lik}$	K	AIC	Δ AIC	w
m1	226.56	2	230.56	8.58	0.01
m2	220.44	3	226.45	4.46	0.06
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- ▶ Model 4 was best supported by the data suggesting that plant height increased with increasing $[\text{CO}_2]$ up to a maximum above which it started to decline.
- ▶ However, model 3 was nearly as well supported as model 4 (evidence ratio = $\frac{0.54}{0.40} = 1.35$).
- ▶ Model 1 was poorly supported showing that it is highly unlikely that $[\text{CO}_2]$ had no effect on plant height.

Careful when interpreting P values after model selection

- ▶ Don't mix model selection and P values

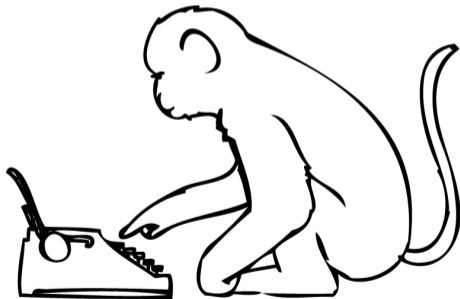
```
##  
## Call:  
## lm(formula = Total.stem.length ~ C02 + I(C02^2))  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max   
## -40.040 -13.688  -4.692  16.462  35.776   
##  
## Coefficients:  
##              Estimate Std. Error t value Pr(>|t|)   
## (Intercept) -4.198e+00  2.022e+01  -0.208  0.83754   
## C02          2.483e-01  8.096e-02   3.068  0.00584 **   
## I(C02^2)    -1.714e-04  6.733e-05  -2.546  0.01881 *   
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##
```

Why not fit “all possible” models and see which one comes out best?

- ▶ leads to fitting of lots of models

e.g. with 10 covariates there are $2^{10} = 1024$ possible regression models (ignoring interactions and polynomial effects)

- ▶ overfitting guaranteed



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Why not use step-wise model selection?

- ▶ leads to fitting of lots of models
- ▶ overfitting guaranteed
- ▶ spurious results guaranteed
- ▶ different procedures don't lead to the same result
- ▶ misleading if explanatory variables are correlated

→ don't do it!

Types of uncertainty

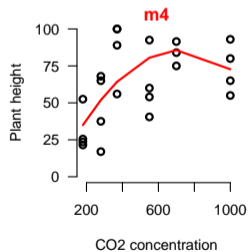
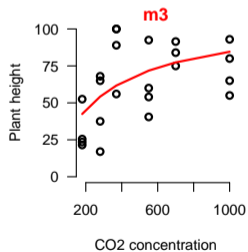
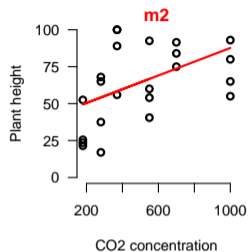
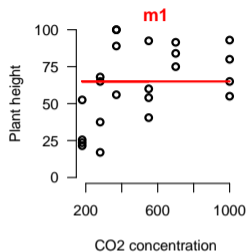
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- ▶ Structural uncertainty (which model is correct)
→ Akaike weights
- ▶ Uncertainty conditional on model structure →
standard errors



"Acacia karroo, bloeityd, Roodeplaat NR" by JMK - Own work. Licensed under CC BY-SA 3.0 via Wikimedia Commons

Model-averaged predictions



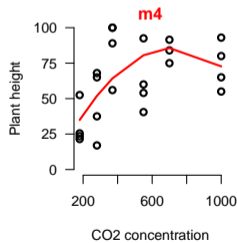
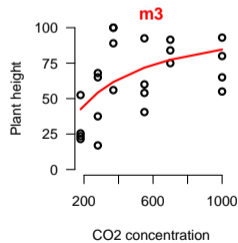
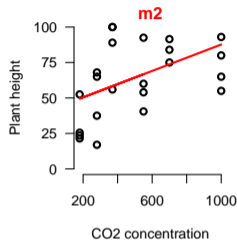
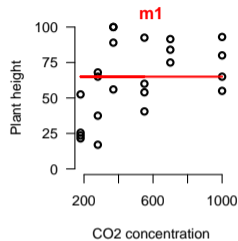
Average?

$$\hat{\theta} = \frac{\sum_{i=1}^R \hat{\theta}_i}{R} = \sum_{i=1}^R \frac{1}{R} \hat{\theta}_i$$

Weighted average using Akaike weights:

$$\hat{\theta} = \sum_{i=1}^R w_i \hat{\theta}_i$$

Model-averaged predictions



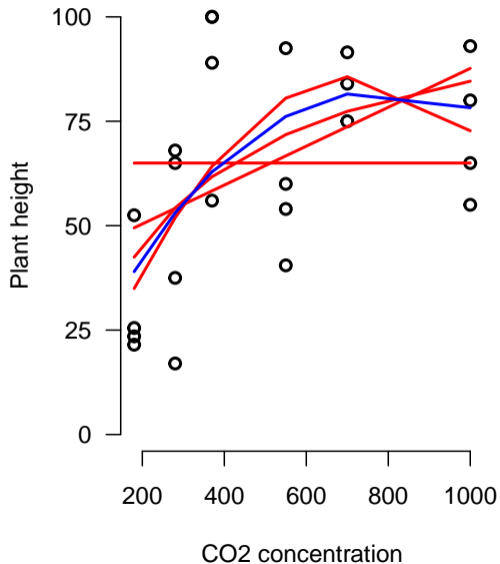
```
cbind(pred200, wi, wi * pred200)
```

```
##      pred200      wi
## [1,]    65.0 0.00739  0.481
## [2,]    50.4 0.05798  2.923
## [3,]    45.2 0.39641 17.925
## [4,]    38.6 0.53822 20.783
```

```
sum(wi * pred200)
```

```
## [1] 42.1
```


Model-averaged predictions



Predictions from individual models
Model-averaged predictions

Unconditional standard error

Averaged point estimate:

$$\hat{\theta} = \sum_{i=1}^R w_i \hat{\theta}_i$$

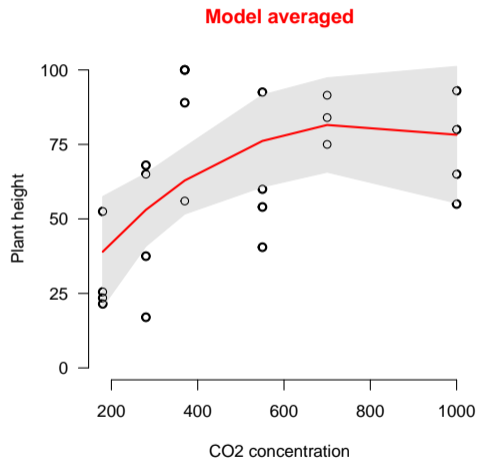
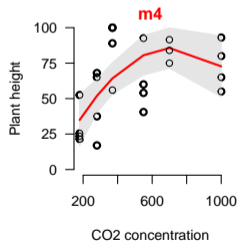
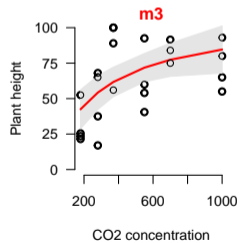
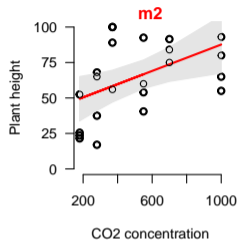
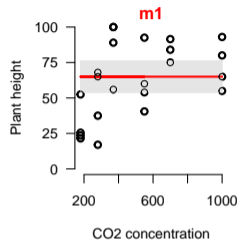
Unconditional measure of uncertainty:

$$\hat{se}(\hat{\theta}) = \sum_{i=1}^R w_i [\hat{var}(\hat{\theta}_i | g_i) + (\hat{\theta}_i - \hat{\theta})^2]^{\frac{1}{2}}$$

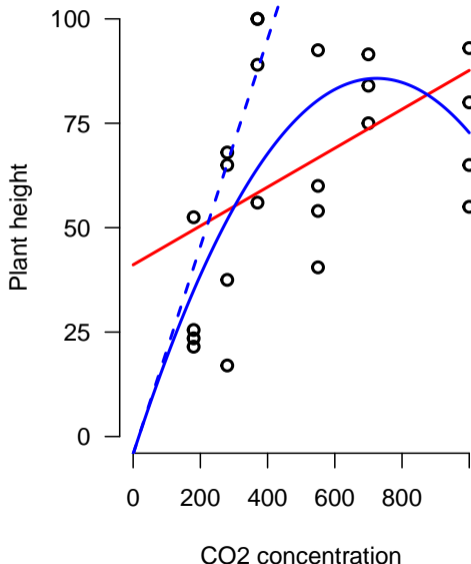
$\hat{var}(\hat{\theta}_i | g_i)$: variance of the model-specific estimate

$(\hat{\theta}_i - \hat{\theta})^2$: variance among model-specific point estimates

Unconditional measures of uncertainty



Averaging parameters? – Careful!



Predictions from model m2

$$y = \beta_0 + \beta_1 \times x$$

Predictions from model m4

$$y = \beta_0 + \beta_1 \times x + \beta_2 \times x^2$$

β_0 and β_0 do not have the same meaning

β_1 and β_1 do not have the same meaning

Literature

Standard textbook on multi-model inference:

- ▶ *Burnham, K. P., and D. R. Anderson. 2002. Model selection and multimodel inference: a practical information-theoretic approach. 2nd edition. Springer, New York.*



Model averaging:

- ▶ *Cade, B. S. 2015. Model averaging and muddled multimodal inferences. Ecology 96:2370–2382. Why you shouldn't average parameters.*
- ▶ *Dormann, C. F., et al 2018. Model averaging in ecology: a review of Bayesian, information-theoretic, and tactical approaches for predictive inference. Ecological Monographs 88:485–504.*